

When are linear predictions of random fields using wrong mean and covariance functions asymptotically optimal?

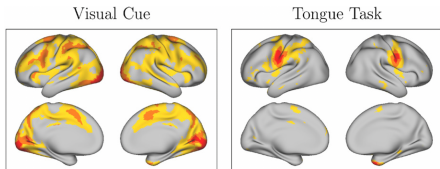
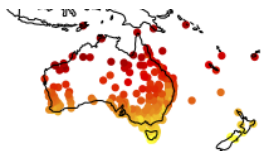
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Linear spatial prediction



Consider a random field $\{Z(x) : x \in \mathcal{X}\}$ on a compact metric space \mathcal{X} .

Aim: Predict its value $Z(x^*)$ at $x^* \in \mathcal{X}$ based on a set of observations $\{Z(x_j)\}_{j=1}^n$ for locations $x_1, \dots, x_n \in \mathcal{X}$ all distinct from x^* .

The **kriging predictor** is the linear predictor

$$Z_n(x^*) = \alpha_0 + \sum_{j=1}^n \alpha_j Z(x_j)$$

based on the observations, where $\alpha_0, \dots, \alpha_n \in \mathbb{R}$ are chosen such that the variance of the error $Z_n(x^*) - Z(x^*)$ is minimized.

Model misspecification

If $m(\cdot)$ and $\varrho(\cdot, \cdot)$ are the mean and the covariance function of Z , then

$$Z_n(x^*) = m(x^*) + \mathbf{c}^\top \boldsymbol{\Sigma}^{-1}(\mathbf{Z} - \mathbf{m}), \quad (*)$$

where

$$\begin{aligned} \mathbf{c} &:= (\varrho(x^*, x_1), \dots, \varrho(x^*, x_n))^\top, & \boldsymbol{\Sigma}_{ij} &:= \varrho(x_i, x_j), \\ \mathbf{Z} &:= (Z(x_1), \dots, Z(x_n))^\top, & \mathbf{m} &:= (m(x_1), \dots, m(x_n))^\top. \end{aligned}$$

\Rightarrow the *kriging predictor* depends only on (m, ϱ) .

Therefore, from now on we assume that $Z \sim \mu = N(m, \varrho)$.

We are interested in the asymptotic behavior of

$$\frac{E[(\tilde{Z}_n(x^*) - Z(x^*))^2]}{E[(Z_n(x^*) - Z(x^*))^2]} \quad \text{as} \quad n \rightarrow \infty,$$

where the linear predictor $\tilde{Z}_n(x^*)$ is computed using $(*)$ with $(\tilde{m}, \tilde{\varrho})$.

What was known?

- M. L. Stein¹ showed that the best linear predictor based on $(\tilde{m}, \tilde{\varrho})$ is asymptotically optimal, as $n \rightarrow \infty$, provided that

$$\mu = N(m, \varrho) \quad \text{and} \quad \tilde{\mu} = N(\tilde{m}, \tilde{\varrho}) \quad \text{are equivalent.}$$

- This result in fact holds *uniformly* with respect to x^* and, moreover, uniformly for each linear functional φ such that $\varphi(Z)$ has finite variance².
- Less restrictive conditions have been derived for some specific cases, such as *periodic* fields³ on $[0, 1]^d$ and *stationary* fields⁴ on \mathbb{R}^d , i.e. $\varrho(x, y) = \varrho(x - y)$, observed on a lattice.

¹M. L. Stein (1988). "Asymptotically efficient prediction of a random field with a misspecified covariance function". In: *Ann. Stat.* 16.1, pp. 55–63

²M. L. Stein (1990). "Uniform asymptotic optimality of linear predictions of a random field using an incorrect second-order structure". In: *Ann. Stat.* 18.2, pp. 850–872

³M. L. Stein (1997). "Efficiency of linear predictors for periodic processes using an incorrect covariance function". In: *J. Statist. Plann. Inference* 58.2, pp. 321–331

⁴M. L. Stein (1999). "Predicting random fields with increasing dense observations". In: *Ann. Appl. Probab.* 9.1, pp. 242–273

Goals of our work

For any constant $c \in (0, \infty)$, the linear predictor based on $(m, c\rho)$ is equal to that based on (m, ρ) , whereas

$$\forall c \neq 1: \quad \mu = \mathbf{N}(m, \rho) \perp \tilde{\mu} = \mathbf{N}(m, c\rho).$$

⇒ Equivalence of the measures μ and $\tilde{\mu}$ is a *sufficient, but not necessary* assumption for asymptotic optimality of linear predictions.

Topics of this talk

- 1 *necessary and sufficient* conditions on $(\tilde{m}, \tilde{\rho})$ for uniform asymptotic optimality of linear predictions;
- 2 explicit conditions for a large class of non-stationary models and
 - equivalence of Gaussian measures;
 - asymptotic optimality of linear predictions.

Outline

① Necessary and sufficient conditions for asymptotic optimality

Linear prediction on compact metric spaces

The assumptions and main result

② Applications

Weakly stationary random fields on \mathbb{R}^d

Common eigenbasis

Generalized Whittle–Matérn fields

Isomorphic Cameron–Martin spaces

Equivalence of measures and asymptotically optimal linear prediction

Summary and simulations

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Setting

- $Z : \mathcal{X} \times \Omega \rightarrow \mathbb{R}$ is a square-integrable Gaussian stochastic process indexed by a connected, compact metric space $(\mathcal{X}, d_{\mathcal{X}})$, with strictly positive and finite Borel measure $\nu_{\mathcal{X}} : \mathcal{B}(\mathcal{X}) \rightarrow [0, \infty)$.

Notation : $L_2 := L_2(\mathcal{X}, \mathcal{B}(\mathcal{X}), \nu_{\mathcal{X}})$.

- Z has mean $m \in L_2$, strictly positive definite and continuous covariance function $\varrho : \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}$, and covariance operator

$$\mathcal{C} : L_2 \rightarrow L_2, \quad (\mathcal{C}w)(x) := \int_{\mathcal{X}} \varrho(x, x') w(x') d\nu_{\mathcal{X}}(x').$$

\mathcal{C} is self-adjoint, positive definite, and trace-class on L_2 .

- We write $\mu = \mathbf{N}(m, \mathcal{C})$ for the Gaussian measure on L_2 induced by the process Z , i.e., for every Borel set $A \in \mathcal{B}(L_2)$,

$$\mu(A) = \mathbb{P}(\{\omega \in \Omega : Z(\cdot, \omega) \in A\}).$$

Kriging predictor as orthogonal projection

- $Z^0 := Z - m$ is a *centered* Gaussian process.
- The vector space $\mathcal{Z}^0 \subset L_2(\Omega, \mathbb{P})$ contains finite linear combinations

$$\mathcal{Z}^0 := \left\{ \sum_{j=1}^K \alpha_j Z^0(x_j) : K \in \mathbb{N}, \alpha_j \in \mathbb{R}, x_j \in \mathcal{X} \right\}, \quad \mathcal{H}^0 := \overline{\mathcal{Z}^0}^{\|\cdot\|_{L_2(\Omega, \mathbb{P})}}.$$

- Every $h = \sum_{j \leq K} \alpha_j Z(x_j)$ has a representation

$$h = c + h^0, \quad \text{with } c \in \mathbb{R}, \quad h^0 \in \mathcal{Z}^0 \subset \mathcal{H}^0.$$

We thus define the Hilbert space $\mathcal{H} := \mathbb{R} \oplus \mathcal{H}^0$,

$$(g, h)_{\mathcal{H}} = \mathbb{E}[g]\mathbb{E}[h] + (g - \mathbb{E}[g], h - \mathbb{E}[h])_{\mathcal{H}^0} = \mathbb{E}[gh].$$

- The kriging predictor h_n of $h \in \mathcal{H}$ based on $\mathcal{H}_n := \mathbb{R} \oplus \mathcal{H}_n^0 \subset \mathbb{R} \oplus \mathcal{H}^0$ is the \mathcal{H} -orthogonal projection of h onto \mathcal{H}_n :

$$h_n \in \mathcal{H}_n : \quad (h_n - h, g_n)_{\mathcal{H}} = \mathbb{E}[(h_n - h)g_n] = 0 \quad \forall g_n \in \mathcal{H}_n,$$

$$h_n \in \mathcal{H}_n : \quad \mathbb{E}[(h_n - h)^2] = \inf_{g_n \in \mathcal{H}_n} \mathbb{E}[(g_n - h)^2].$$

Problem formulation

$$h_n \in \mathcal{H}_n : \quad (h_n - h, g_n)_{\mathcal{H}} = \mathbb{E}[(h_n - h)g_n] = 0 \quad \forall g_n \in \mathcal{H}_n,$$

Main question

What happens if, instead of h_n , we use the linear predictor \tilde{h}_n which is the kriging predictor if $\tilde{\mu} = \mathbf{N}(\tilde{m}, \tilde{C})$ was the correct model?

“Computing orthogonal projections with the wrong inner product”

Consistent kriging prediction

- We require that, for every $h \in \mathcal{H}$, the corresponding kriging predictors $\{h_n\}_{n \in \mathbb{N}}$ are consistent in the sense that

$$\lim_{n \rightarrow \infty} \mathbb{E}[(h_n - h)^2] = \lim_{n \rightarrow \infty} \|h_n - h\|_{\mathcal{H}}^2 = 0. \quad (\text{density})$$

- Let \mathcal{S}_{adm}^{μ} contain all *admissible* sequences $\{\mathcal{H}_n\}_{n \in \mathbb{N}}$ of subspaces $\mathcal{H}_n \subset \mathcal{H}$ which provide μ -consistent kriging prediction,

$$\mathcal{S}_{adm}^{\mu} := \left\{ \{\mathcal{H}_n\}_{n \in \mathbb{N}} \mid \forall n \in \mathbb{N} : \mathcal{H}_n = \mathbb{R} \oplus \mathcal{H}_n^0 \text{ with } \dim(\mathcal{H}_n^0) = n, \right. \\ \left. \forall h \in \mathcal{H} : \{h_n\}_{n \in \mathbb{N}} \text{ satisfy } (\text{density}) \right\}.$$

Example

Suppose that m, ϱ are continuous and that every \mathcal{H}_n is generated by point observations $Z(x_1), Z(x_2), \dots$. Then, $\{\mathcal{H}_n\}_{n \in \mathbb{N}} \in \mathcal{S}_{adm}^{\mu}$ if $\{x_j\}_{j \in \mathbb{N}}$ is a sequence in $(\mathcal{X}, d_{\mathcal{X}})$ which *accumulates at any* $x^* \in \mathcal{X}$.

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Main assumptions

The Assumptions

Let $\varrho, \tilde{\varrho}: \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}$ be two continuous, (strictly) positive definite covariance functions. Assume that the corresponding covariance operators $\mathcal{C}, \tilde{\mathcal{C}}: L_2 \rightarrow L_2$, and $m, \tilde{m} \in L_2$ are such that:

- I. The vector spaces

$$\mathcal{C}^{1/2}(L_2), (\mathcal{C}^{-1/2} \cdot, \mathcal{C}^{-1/2} \cdot)_{L_2} \quad \text{and} \quad \tilde{\mathcal{C}}^{1/2}(L_2), (\tilde{\mathcal{C}}^{-1/2} \cdot, \tilde{\mathcal{C}}^{-1/2} \cdot)_{L_2}$$

are *norm equivalent* Hilbert spaces.

- II. The difference of the means satisfies $m - \tilde{m} \in \mathcal{C}^{1/2}(L_2)$.
- III. There exists a positive constant $a \in (0, \infty)$ such that the operator

$$T_a: L_2 \rightarrow L_2, \quad T_a := \mathcal{C}^{-1/2} \tilde{\mathcal{C}} \mathcal{C}^{-1/2} - a\mathcal{I}$$

is *compact* on L_2 . Here \mathcal{I} denotes the identity on L_2 .

Theorem (Asymptotic optimality, K. and Bolin, 2020)

Let $\mu = N(m, \mathcal{C})$ and $\tilde{\mu} = N(\tilde{m}, \tilde{\mathcal{C}})$. Let h_n, \tilde{h}_n denote the best linear predictors of h based on \mathcal{H}_n and μ resp. $\tilde{\mu}$. Then, any of the assertions,

$$\lim_{n \rightarrow \infty} \sup_{h \in \mathcal{H}_{-n}} \frac{E[(\tilde{h}_n - h)^2]}{E[(h_n - h)^2]} = 1,$$

$$\lim_{n \rightarrow \infty} \sup_{h \in \mathcal{H}_{-n}} \frac{\tilde{E}[(h_n - h)^2]}{\tilde{E}[(\tilde{h}_n - h)^2]} = 1,$$

$$\lim_{n \rightarrow \infty} \sup_{h \in \mathcal{H}_{-n}} \left| \frac{\tilde{E}[(h_n - h)^2]}{E[(h_n - h)^2]} - a \right| = 0,$$

$$\lim_{n \rightarrow \infty} \sup_{h \in \mathcal{H}_{-n}} \left| \frac{E[(\tilde{h}_n - h)^2]}{\tilde{E}[(\tilde{h}_n - h)^2]} - \frac{1}{a} \right| = 0,$$

holds for all $\{\mathcal{H}_n\}_{n \in \mathbb{N}} \in \mathcal{S}_{adm}^\mu$ if and only if the Assumptions I–III are satisfied. The constant $a \in (0, \infty)$ is the same as that in Assumption III.

Comparison with the Feldman–Hájek theorem

Equivalence of Gaussian measures

Let μ and $\tilde{\mu}$ be two measures on $(L_2, \mathcal{B}(L_2))$. Then, μ and $\tilde{\mu}$ are called

- equivalent if $\mu(A) = 0 \Leftrightarrow \tilde{\mu}(A) = 0$;
- orthogonal if there exists $B \in \mathcal{B}(L_2)$ with $\mu(B) = 0$ and $\tilde{\mu}(B) = 1$.

Theorem (Feldman–Hájek)

Two Gaussian measures $\mu = N(m, C)$ and $\tilde{\mu} = N(\tilde{m}, \tilde{C})$ on a separable Hilbert space $(E, (\cdot, \cdot)_E)$ are either orthogonal or equivalent. They are equivalent if and only if the following conditions are satisfied:

- The Cameron–Martin spaces $C^{1/2}(E)$, $\tilde{C}^{1/2}(E)$ are norm equivalent.
- The difference of the means satisfies $m - \tilde{m} \in C^{1/2}(E)$.
- The operator $T_1 := C^{-1/2}\tilde{C}C^{-1/2} - \text{Id}_E$ is Hilbert–Schmidt on E .

Proposition (Role of Assumption I)

Let $\mu = \mathbf{N}(m, \mathcal{C})$, $\tilde{\mu} = \mathbf{N}(\tilde{m}, \tilde{\mathcal{C}})$, and define $\mathcal{H}^0, \tilde{\mathcal{H}}^0$ with respect to the measures μ and $\tilde{\mu}$, respectively. The following are equivalent:

- (i) Assumption I is satisfied.
- (ii) The linear operator $\tilde{\mathcal{C}}^{-1/2} \mathcal{C}^{-1/2} : L_2 \rightarrow L_2$ is an isomorphism.
- (iii) The Hilbert spaces $\mathcal{H}^0, \tilde{\mathcal{H}}^0$ are norm equivalent. In particular,

$$\exists k_0, k_1 > 0 : \quad k_0 \text{Var}[h] \leq \widetilde{\text{Var}}[h] \leq k_1 \text{Var}[h] \quad \forall h \in \mathcal{H}.$$

- (iv) There exist $0 < k \leq K < \infty$ such that, for every $\{\mathcal{H}_n\}_{n \in \mathbb{N}} \in \mathcal{S}_{adm}^\mu$, for all $n \in \mathbb{N}$, and every $h \in \mathcal{H}_{-n}$,

$$\frac{\widetilde{\text{Var}}[h_n - h]}{\text{Var}[h_n - h]}, \frac{\text{Var}[\tilde{h}_n - h]}{\widetilde{\text{Var}}[\tilde{h}_n - h]}, \frac{\text{Var}[\tilde{h}_n - h]}{\text{Var}[h_n - h]}, \frac{\widetilde{\text{Var}}[h_n - h]}{\widetilde{\text{Var}}[\tilde{h}_n - h]} \in [k, K].$$

h_n, \tilde{h}_n are the best linear predictors of h based on \mathcal{H}_n and μ resp. $\tilde{\mu}$.

Proposition (Role of Assumptions I and III)

Let $\mu = N(m, \mathcal{C})$ and $\tilde{\mu} = N(\tilde{m}, \tilde{\mathcal{C}})$. Let h_n, \tilde{h}_n denote the best linear predictors of h based on \mathcal{H}_n and μ resp. $\tilde{\mu}$. Then, any of the assertions,

$$\lim_{n \rightarrow \infty} \sup_{h \in \mathcal{H}_{-n}} \frac{\text{Var}[\tilde{h}_n - h]}{\text{Var}[h_n - h]} = 1,$$

$$\lim_{n \rightarrow \infty} \sup_{h \in \mathcal{H}_{-n}} \frac{\widetilde{\text{Var}}[h_n - h]}{\widetilde{\text{Var}}[\tilde{h}_n - h]} = 1,$$

$$\lim_{n \rightarrow \infty} \sup_{h \in \mathcal{H}_{-n}} \left| \frac{\widetilde{\text{Var}}[h_n - h]}{\text{Var}[h_n - h]} - a \right| = 0,$$

$$\lim_{n \rightarrow \infty} \sup_{h \in \mathcal{H}_{-n}} \left| \frac{\text{Var}[\tilde{h}_n - h]}{\widetilde{\text{Var}}[\tilde{h}_n - h]} - \frac{1}{a} \right| = 0,$$

holds for all $\{\mathcal{H}_n\}_{n \in \mathbb{N}} \in \mathcal{S}_{adm}^\mu$ if and only if Assumptions I and III are fulfilled. The constant $a \in (0, \infty)$ is the same as that in Assumption III.

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Examples of valid covariance models and metric spaces

Valid covariance functions are given by the Matérn class,

$$\varrho(x, x') := \varrho_0(d_{\mathcal{X}}(x, x')), \quad \varrho_0(r) = \frac{\sigma^2}{2^{\nu-1}\Gamma(\nu)} (\kappa r)^{\nu} K_{\nu}(\kappa r), \quad r \geq 0,$$

on the compact metric space $(\mathcal{X}, d_{\mathcal{X}})$, where

- $\mathcal{X} \subset \mathbb{R}^d$ is a *connected, compact Euclidean domain* for $\nu, \kappa, \sigma^2 > 0$.
- $\mathcal{X} := \mathbb{S}^d$ is the *d-sphere* equipped with the great circle distance, for every $\nu \in (0, 1/2]$ and all $\kappa, \sigma^2 > 0$ ⁵.
- \mathcal{X} is a *graph with Euclidean edges* equipped with the resistance metric for every for every $\nu \in (0, 1/2]$ and all $\kappa, \sigma^2 > 0$ ⁶.

⁵T. Gneiting (2013). "Strictly and non-strictly positive definite functions on spheres". In: *Bernoulli* 19.4, pp. 1327–1349

⁶E. Anderes, J. Møller, and J. G. Rasmussen (2020). "Isotropic covariance functions on graphs and their edges". In: *Ann. Statist.* 48.4, pp. 2478–2503

Weakly stationary random fields on \mathbb{R}^d

Setting: $\mathcal{X} \subset \mathbb{R}^d$ and $\varrho|_{\mathcal{X} \times \mathcal{X}}$, $\tilde{\varrho}|_{\mathcal{X} \times \mathcal{X}}$ are restrictions of continuous, positive definite, translation invariant functions $\varrho, \tilde{\varrho}: \mathbb{R}^d \times \mathbb{R}^d \rightarrow \mathbb{R}$.

$$\exists \varrho_0, \tilde{\varrho}_0: \mathbb{R}^d \rightarrow \mathbb{R} \text{ even: } \quad \varrho(x, x') = \varrho_0(x - x'), \quad \tilde{\varrho}(x, x') = \tilde{\varrho}_0(x - x').$$

The spectral density f and ϱ_0 relate via the inversion formula:

$$\forall \omega \in \mathbb{R}^d: \quad f(\omega) = \frac{1}{(2\pi)^d} (\mathcal{F}\varrho_0)(\omega), \quad (\mathcal{F}\varrho_0)(\omega) := \int_{\mathbb{R}^d} e^{-i\omega \cdot x} \varrho_0(x) dx.$$

We define $\mathcal{F}_{\mathcal{X}} := \mathcal{F} \circ E_{\mathcal{X}}^0$, where $E_{\mathcal{X}}^0$ is the zero extension, and

$$\mathcal{F}_{\mathcal{X}}(L_2(\mathcal{X})) = \{ \hat{w}: \mathbb{R}^d \rightarrow \mathbb{C} \mid \exists w \in L_2(\mathcal{X}) : \hat{w} = \mathcal{F}_{\mathcal{X}} w \} \subset L_2(\mathbb{R}^d; \mathbb{C}).$$

The Hilbert space H_f (over \mathbb{R}) is the closure of $\mathcal{F}_{\mathcal{X}}(L_2(\mathcal{X}))$ with respect to norm induced by the inner product

$$(\hat{v}_1, \hat{v}_2)_{H_f} := \int_{\mathbb{R}^d} f(\omega) \hat{v}_1(\omega) \overline{\hat{v}_2(\omega)} d\omega, \quad H_f := \overline{\mathcal{F}_{\mathcal{X}}(L_2(\mathcal{X}))}^{\|\cdot\|_{H_f}}.$$

Proposition (Assumptions I and III in terms of spectral densities)

Suppose that $\mathcal{C}, \tilde{\mathcal{C}} : L_2(\mathcal{X}) \rightarrow L_2(\mathcal{X})$ pertain to restrictions (to $\mathcal{X} \times \mathcal{X}$) of translation invariant covariance functions $\varrho, \tilde{\varrho} : \mathbb{R}^d \times \mathbb{R}^d \rightarrow \mathbb{R}$, which have spectral densities $f, \tilde{f} : \mathbb{R}^d \rightarrow [0, \infty)$.

Then, Assumptions I and III are satisfied if and only if:

- I' The spaces H_f and $H_{\tilde{f}}$ are isomorphic with equivalent norms, i.e., there exist constants $0 < k \leq K < \infty$ such that

$$k \|\hat{v}\|_{H_f}^2 \leq \int_{\mathbb{R}^d} \tilde{f}(\omega) |\hat{v}(\omega)|^2 d\omega \leq K \|\hat{v}\|_{H_f}^2 \quad \forall \hat{v} \in \mathcal{F}_{\mathcal{X}}(L_2(\mathcal{X})).$$

- III' There exists $a \in (0, \infty)$ such that the linear operator $\hat{T}_a := S - a\mathcal{I}_{H_f}$ is compact on H_f , where \mathcal{I}_{H_f} denotes the identity on H_f and $S : H_f \rightarrow H_f$ is defined by

$$(S\hat{v}_1, \hat{v}_2)_{H_f} = \int_{\mathbb{R}^d} \tilde{f}(\omega) \hat{v}_1(\omega) \overline{\hat{v}_2(\omega)} d\omega \quad \forall \hat{v}_1, \hat{v}_2 \in H_f.$$

$$\varrho(x, x') := \varrho_0(\|x - x'\|_{\mathbb{R}^d}), \quad \varrho_0(r) = \frac{\sigma^2}{2^{\nu-1}\Gamma(\nu)} (\kappa r)^\nu K_\nu(\kappa r), \quad r \geq 0,$$

Parameters: $\nu, \kappa, \sigma^2 > 0$.

Example (Matérn covariance family)

Assumptions I and III are satisfied if and only if $\nu = \tilde{\nu}$. In this case:

$$a = \frac{\tilde{\sigma}^2 \tilde{\kappa}^{2\nu}}{\sigma^2 \kappa^{2\nu}} \in (0, \infty).$$

For equivalence of the corresponding Gaussian measures, $a = 1$ is necessary. Indeed, Zhang⁷ and Anderes⁸ showed that, for $\nu = \tilde{\nu}$,

$$\mu \sim \tilde{\mu} \iff \begin{cases} \sigma^2 \kappa^{2\nu} = \tilde{\sigma}^2 \tilde{\kappa}^{2\nu} & \text{for } d \leq 3, \\ \kappa = \tilde{\kappa} \text{ and } \sigma^2 = \tilde{\sigma}^2 & \text{for } d \geq 5. \end{cases}$$

⁷H. Zhang (2004). "Inconsistent estimation and asymptotically equal interpolations in model-based geostatistics". In: *J. Amer. Statist. Assoc.* 99.465, pp. 250–261

⁸E. Anderes (2010). "On the consistent separation of scale and variance for Gaussian random fields". In: *Ann. Statist.* 38.2, pp. 870–893

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Covariance operators with the same eigenbasis

Corollary

Suppose that $\mathcal{C}, \tilde{\mathcal{C}}$ are self-adjoint, positive definite, trace-class operators on $L_2(\mathcal{X}, \nu_{\mathcal{X}})$ which diagonalize with respect to the same orthonormal basis $\{e_j\}_{j \in \mathbb{N}}$ for $L_2(\mathcal{X}, \nu_{\mathcal{X}})$, i.e., there are $\{\gamma_j\}_{j \in \mathbb{N}}, \{\tilde{\gamma}_j\}_{j \in \mathbb{N}} \subset (0, \infty)$ s.t.

$$\mathcal{C}e_j = \gamma_j e_j \quad \text{and} \quad \tilde{\mathcal{C}}e_j = \tilde{\gamma}_j e_j \quad \forall j \in \mathbb{N}.$$

Consider $\mu := N(0, \mathcal{C})$ and $\tilde{\mu} := N(0, \tilde{\mathcal{C}})$.

- The Cameron–Martin spaces for μ and $\tilde{\mu}$ are isomorphic if and only if there exist $c_-, c_+ \in \mathbb{R}_+$ such that $\tilde{\gamma}_j/\gamma_j \in [c_-, c_+]$ for all $j \in \mathbb{N}$.
- μ and $\tilde{\mu}$ are equivalent if and only if $\sum_{j \in \mathbb{N}} (\tilde{\gamma}_j/\gamma_j - 1)^2 < \infty$.
- Assumptions I and III are satisfied if and only if there exists a constant $a \in (0, \infty)$ such that $\lim_{j \rightarrow \infty} \tilde{\gamma}_j/\gamma_j = a$.

Consider the shifted negative Dirichlet Laplacian:

$$Lv := (-\Delta + \kappa^2)v, \quad v \in \mathcal{D}(L) := H^2(\mathcal{D}) \cap H_0^1(\mathcal{D}),$$

where $\mathcal{D} \subsetneq \mathbb{R}^d$ is a bounded open Lipschitz domain.

Corollary (“Classical” Whittle–Matérn fields)

Let $d \in \mathbb{N}$, $\beta, \tilde{\beta} > d/4$, $\tau, \tilde{\tau} > 0$, and let L, \tilde{L} have shift parameters $\kappa^2 \geq 0$ and $\tilde{\kappa}^2 \geq 0$, respectively. Consider on $L_2(\mathcal{D})$ the Gaussian measures

$$\mu = N(0, \tau^{-2}L^{-2\beta}) \quad \text{and} \quad \tilde{\mu} = N(0, \tilde{\tau}^{-2}\tilde{L}^{-2\tilde{\beta}}).$$

- The Cameron–Martin spaces for μ and $\tilde{\mu}$ are isomorphic, with equivalent norms, if and only if $\beta = \tilde{\beta}$.
- μ and $\tilde{\mu}$ are equivalent if and only if

$$\begin{cases} \beta = \tilde{\beta} \text{ and } \tau = \tilde{\tau} & \text{if } d \leq 3, \\ \beta = \tilde{\beta}, \tau = \tilde{\tau} \text{ and } \kappa^2 = \tilde{\kappa}^2 & \text{if } d \geq 4. \end{cases}$$

- For any $d \in \mathbb{N}$, Assumptions I & III are fulfilled if and only if $\beta = \tilde{\beta}$.

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Generalized Whittle–Matérn fields on $\mathcal{D} \not\subseteq \mathbb{R}^d$

Next, we consider Gaussian measures on $L_2(\mathcal{D})$, $\mathcal{D} \not\subseteq \mathbb{R}^d$, of the form

$$\mu_d(m; \beta, \mathbf{a}, \kappa) := \mathbf{N}(m, L^{-2\beta}),$$

where $m \in L_2(\mathcal{D})$, $\beta > d/4$ and

$$Lv = -\nabla \cdot (\mathbf{a} \nabla v) + \kappa^2 v, \quad v \in \mathcal{D}(L) \subseteq L_2(\mathcal{D}) \cap H_0^1(\mathcal{D}).$$

We suppose that \mathbf{a} and κ and the domain $\mathcal{D} \subset \mathbb{R}^d$ satisfy the following.

Setting WM

i. $\mathbf{a} : \bar{\mathcal{D}} \rightarrow \mathbb{R}^{d \times d}$ is symmetric and uniformly positive definite, i.e.,

$$\exists a_0 > 0 : \quad \forall \xi \in \mathbb{R}^d : \quad \text{ess inf}_{s \in \mathcal{D}} \xi^\top \mathbf{a}(s) \xi \geq a_0 \|\xi\|_{\mathbb{R}^d}^2,$$

and $\mathbf{a} = (\mathbf{a}_{jk})_{j,k=1}^d$ is smooth, $\mathbf{a}_{jk} \in C^\infty(\bar{\mathcal{D}})$ for all $j, k \in \{1, \dots, d\}$.

ii. $\kappa : \bar{\mathcal{D}} \rightarrow \mathbb{R}$ is smooth, $\kappa \in C^\infty(\bar{\mathcal{D}})$.

iii. The domain $\mathcal{D} \subset \mathbb{R}^d$ has a smooth boundary $\partial\mathcal{D}$ of class C^∞ .

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Lemma (Cameron–Martin space for $\mu_d(m; \beta, \mathbf{a}, \kappa)$)

Suppose Setting WM. For every $\beta > d/4$, the Cameron–Martin space of the Gaussian measure $\mu_d(m; \beta, \mathbf{a}, \kappa)$ is

$$\mathcal{C}^{1/2}(L_2(\mathcal{D})) = \dot{H}_L^{2\beta} := \mathcal{D}(L^\beta) \subseteq L_2(\mathcal{D}), \quad \|v\|_{2\beta, L} := \|L^\beta v\|_{L_2(\mathcal{D})}.$$

It is a subspace of the Sobolev space $H^{2\beta}(\mathcal{D})$ and

$$(\dot{H}_L^{2\beta}, \|\cdot\|_{2\beta, L}) \hookrightarrow (H^{2\beta}(\mathcal{D}), \|\cdot\|_{H^{2\beta}(\mathcal{D})}) \hookrightarrow (C^0(\overline{\mathcal{D}}), \|\cdot\|_{C^0(\overline{\mathcal{D}})})$$

Furthermore, provided that $2\beta \notin \mathcal{E}$, where

$$\mathcal{E} := \{2k + 1/2 : k \in \mathbb{N}_0\},$$

we have the identification

$$\dot{H}_L^{2\beta} = \left\{ v \in H^{2\beta}(\mathcal{D}) : (\kappa^2 - \nabla \cdot (\mathbf{a} \nabla))^j v = 0 \text{ in } L_2(\partial\mathcal{D}) \quad \forall j = 0, \dots, \left\lfloor \beta - \frac{1}{4} \right\rfloor \right\},$$

and on $\dot{H}_L^{2\beta}$ the Sobolev norm $\|\cdot\|_{H^{2\beta}(\mathcal{D})}$ and $\|\cdot\|_{2\beta, L}$ are equivalent.

Theorem (Isomorphic Cameron–Martin spaces)

Suppose Setting WM for each of the parameter tuples (\mathbf{a}, κ) , $(\tilde{\mathbf{a}}, \tilde{\kappa})$ and for $\mathcal{D} \not\subseteq \mathbb{R}^d$. Let $\beta, \tilde{\beta} > d/4$ be such that $2\beta \notin \mathcal{E}$.

The Cameron–Martin spaces of two Gaussian measures

$$\mu_d(0; \beta, \mathbf{a}, \kappa) \quad \text{and} \quad \mu_d(0; \tilde{\beta}, \tilde{\mathbf{a}}, \tilde{\kappa})$$

are isomorphic with equivalent norms if and only if $\beta = \tilde{\beta}$ and for every $j \in \mathbb{N}_0$ with $j \leq \lfloor \beta - 5/4 \rfloor$ the following hold:

$$\begin{aligned} \forall v \in \dot{H}_L^{2\beta} : & \quad (\kappa^2 - \nabla \cdot (\mathbf{a}\nabla))^j (\delta_{\kappa^2} - \nabla \cdot (\delta_{\mathbf{a}}\nabla))v = 0 \quad \text{in } L_2(\partial\mathcal{D}), \\ \forall \tilde{v} \in \dot{H}_L^{2\tilde{\beta}} : & \quad (\tilde{\kappa}^2 - \nabla \cdot (\tilde{\mathbf{a}}\nabla))^j (\delta_{\tilde{\kappa}^2} - \nabla \cdot (\delta_{\tilde{\mathbf{a}}}\nabla))\tilde{v} = 0 \quad \text{in } L_2(\partial\mathcal{D}). \end{aligned} \quad (\text{BCs})$$

Here, $\delta_{\kappa^2}(s) := \tilde{\kappa}^2(s) - \kappa^2(s)$ and $\delta_{\mathbf{a}}(s) := \tilde{\mathbf{a}}(s) - \mathbf{a}(s)$ for all $s \in \overline{\mathcal{D}}$.

\Rightarrow The behavior of δ_{κ^2} and $\delta_{\mathbf{a}}$ on the boundary $\partial\mathcal{D}$ matters!

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Conditions on the parameters for equivalent measures

Theorem (Characterizing equivalence of the measures μ and $\tilde{\mu}$)

Suppose Setting WM for each of the parameter tuples (\mathbf{a}, κ) , $(\tilde{\mathbf{a}}, \tilde{\kappa})$ and for $\mathcal{D} \not\subseteq \mathbb{R}^d$. Let $m, \tilde{m} \in L_2(\mathcal{D})$ and $\beta, \tilde{\beta} > d/4$ be such that $2\beta \notin \mathcal{E}$.

- In dimension $d \leq 3$, the Gaussian measures $\mu_d(m; \beta, \mathbf{a}, \kappa)$ and $\mu_d(\tilde{m}; \tilde{\beta}, \tilde{\mathbf{a}}, \tilde{\kappa})$ are equivalent if and only if
 - $\beta = \tilde{\beta}$,
 - the boundary conditions (BCs) hold for every $j = 0, 1, \dots, \lfloor \beta - 5/4 \rfloor$,
 - $m - \tilde{m} \in \dot{H}_L^{2\beta}$,
 - $\mathbf{a} = \tilde{\mathbf{a}}$.
- In dimension $d \geq 4$, the Gaussian measures $\mu_d(m; \beta, \mathbf{a}, \kappa)$ and $\mu_d(\tilde{m}; \tilde{\beta}, \tilde{\mathbf{a}}, \tilde{\kappa})$ are equivalent if and only if
 - $\beta = \tilde{\beta}$,
 - $\kappa^2 = \tilde{\kappa}^2$,
 - $m - \tilde{m} \in \dot{H}_L^{2\beta}$,
 - $\mathbf{a} = \tilde{\mathbf{a}}$.

Conditions on the parameters for optimal linear prediction

Theorem (Uniformly asymptotically optimal linear prediction)

Suppose Setting WM for each of the parameter tuples (\mathbf{a}, κ) , $(\tilde{\mathbf{a}}, \tilde{\kappa})$ and for $\mathcal{D} \not\subseteq \mathbb{R}^d$. Let $m, \tilde{m} \in L_2(\mathcal{D})$ and $\beta, \tilde{\beta} > d/4$ be such that $2\beta \notin \mathcal{E}$.

Consider the Gaussian measures

$$\mu_d(m; \beta, \mathbf{a}, \kappa) \quad \text{and} \quad \mu_d(\tilde{m}; \tilde{\beta}, \tilde{\mathbf{a}}, \tilde{\kappa})$$

Then, the Assumptions I–III (for uniformly asymptotically optimal linear prediction) are satisfied if and only if

- $\beta = \tilde{\beta}$,
- the boundary conditions (BCs) hold for every $j = 0, 1, \dots, \lfloor \beta - 5/4 \rfloor$,
- $m - \tilde{m} \in \dot{H}_L^{2\beta}$,
- here exists a constant $c > 0$ such that $c\mathbf{a} = \tilde{\mathbf{a}}$.

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Summary

Let $\mu := \mu_d(0; \beta, \mathbf{a}, \kappa)$, $\tilde{\mu} := \mu_d(0; \tilde{\beta}, \tilde{\mathbf{a}}, \tilde{\kappa})$ be Gaussian measures for generalized Whittle–Matérn fields with parameters $(\beta, \mathbf{a}, \kappa)$ resp. $(\tilde{\beta}, \tilde{\mathbf{a}}, \tilde{\kappa})$.

Interval for β , assuming that $\beta \notin \{k + 1/4 : k \in \mathbb{N}\}$

| Conditions for | $(d/4, 9/4)$ | $(9/4, 13/4)$ | $(13/4, \infty)$ |
|---|--|---|---|
| Asymptotically optimal linear prediction | $\beta = \tilde{\beta}$, $\mathbf{ca} = \tilde{\mathbf{a}}$ for some $c \in (0, \infty)$ | $\beta = \tilde{\beta}$, $\mathbf{ca} = \tilde{\mathbf{a}}$, $(\mathbf{a}\nabla\delta_{c,\kappa^2}) _{\partial\mathcal{D}} \cdot \mathbf{n} = 0$ | $\beta = \tilde{\beta}$, $\mathbf{ca} = \tilde{\mathbf{a}}$ + b.c. on δ_{c,κ^2} |
| Equivalence of measures in dimension $d \leq 3$ | $\beta = \tilde{\beta}$, $\mathbf{a} = \tilde{\mathbf{a}}$ | $\beta = \tilde{\beta}$, $\mathbf{a} = \tilde{\mathbf{a}}$, $(\mathbf{a}\nabla\delta_{1,\kappa^2}) _{\partial\mathcal{D}} \cdot \mathbf{n} = 0$ | $\beta = \tilde{\beta}$, $\mathbf{a} = \tilde{\mathbf{a}}$ + b.c. on δ_{1,κ^2} |
| Equivalence of measures in dimension $d \geq 4$ | $\beta = \tilde{\beta}$, $\mathbf{a} = \tilde{\mathbf{a}}$, $\kappa^2 = \tilde{\kappa}^2$ | | |

$\delta_{c,\kappa^2}(s) := \tilde{\kappa}^2(s) - c\kappa^2(s)$, and \mathbf{n} is the outward pointing normal on $\partial\mathcal{D}$.

Simulation 1: The difference between κ^2 and \mathbf{a}

$$(-\nabla \cdot (\mathbf{a}\nabla) + \kappa^2)^\beta (\tau Z) = \mathcal{W} \quad \text{in } \mathcal{D} = (0, 1).$$

True model: $\beta = 1$, $\mathbf{a} \equiv 1$, $\kappa^2 \equiv 1200$, and $\tau = \frac{1}{2}\kappa^{-3/2}$.

Two misspecified models: correct values of β, τ , and we set

$$(\kappa^2(s), \mathbf{a}(s)) = \begin{cases} (1200f(s)^{-1}, 1) & \text{for model 1,} \\ (1200, f(s)) & \text{for model 2,} \end{cases} \quad s \in \overline{\mathcal{D}} = [0, 1],$$

where $f(s) := 1 + \frac{1}{2} \operatorname{erf}\left(\frac{\delta(s-0.5)}{\sqrt{2}}\right)$.

As a measure of accuracy, we use

$$\mathcal{E}_n(h) := \frac{\mathbb{E}[(\tilde{h}_n - h)^2]}{\mathbb{E}[(h_n - h)^2]} - 1.$$

We predict $Z(x_0)$ as well as the integrals $l_\ell := (Z, e_\ell)_{L_2(\mathcal{D})}$ and set

$$\mathcal{E}_{l,n}^{\max} := \max\{\mathcal{E}_{l,n}^\ell : n+1 \leq \ell \leq N\}, \quad \mathcal{E}_{l,n}^\ell := \mathcal{E}_n(l_\ell), \quad \ell \in \{n+1, \dots, N\}.$$

Results of Simulation 1

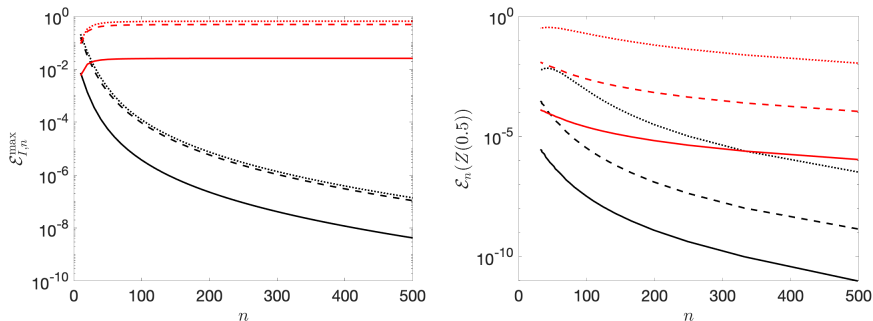


Figure: The results for model 1 (black) and model 2 (red) for the first example with integral observations (left) and point observations (right).

Solid lines correspond to $\delta = 1$, dashed to $\delta = 10$, and dotted to $\delta = 100$.

Simulation 2: The effect of the smoothness parameter β

For $\beta \in \{1, 2, 3\}$, we consider

$$(-\nabla \cdot (\mathbf{a}\nabla) + \kappa^2)^\beta (\tau Z) = \mathcal{W} \quad \text{in } \mathcal{D} = (0, 1).$$

True model: $\mathbf{a} \equiv 1$, $\kappa^2 \equiv 100(4\beta - 1)$, and $\tau > 0$.

Two misspecified models: correct values of β , \mathbf{a} , τ , and we set

$$\kappa^2(s) = 100(4\beta - 1) \cdot \begin{cases} 1 - 1.5s^2 + s^3 & \text{for model 1,} \\ 1 + s - 1.5s^3 & \text{for model 2,} \end{cases} \quad s \in \overline{\mathcal{D}} = [0, 1].$$

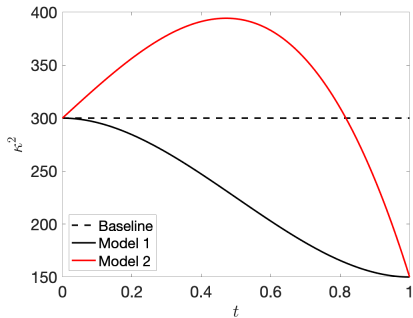
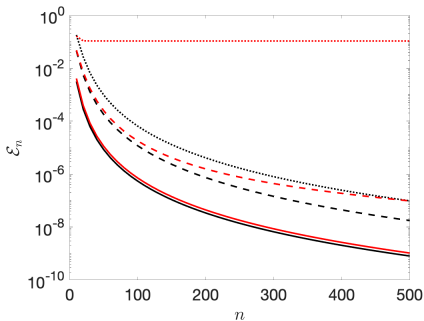
We again predict the integrals $I_\ell := (Z, e_\ell)_{L_2(\mathcal{D})}$ and consider

$$\mathcal{E}_{I,n}^{\max} := \max\{\mathcal{E}_{I,n}^\ell : n+1 \leq \ell \leq N\}, \quad \mathcal{E}_{I,n}^\ell := \mathcal{E}_n(I_\ell), \quad \ell \in \{n+1, \dots, N\},$$

where

$$\mathcal{E}_n(h) := \frac{\mathbb{E}[(\tilde{h}_n - h)^2]}{\mathbb{E}[(h_n - h)^2]} - 1.$$

Results of Simulation 2



Left: The results for model 1 (black) and model 2 (red) in the second example, with $\beta = 1$ (solid), $\beta = 2$ (dashed), and $\beta = 3$ (dotted).

Right: κ^2 for the two models when $\beta = 1$.

References

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- D. Bolin and K. Kirchner (2021). Equivalence of measures and asymptotically optimal linear prediction for Gaussian random fields with fractional-order covariance operators. Preprint: [arXiv:2101.07860](https://arxiv.org/abs/2101.07860).

Thank you for your attention!