

# **On a simple model for ruin of two insurance companies**

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We consider a two-dimensional risk model in which two companies split the amount they pay out of each claim in proportions  $\delta_1$  and  $\delta_2$  where  $\delta_1 + \delta_2 = 1$ , and receive premiums at rates  $c_1$  and  $c_2$ , respectively. Claims arrive according to a renewal process. We analyse the ruin problems of the risk process leaving the positive quadrant and of entering the negative quadrant under a light-tail assumption.

## **The multidimensional renewal risk model.**

In collective risk theory the reserves process  $X$  of an insurance company is modeled as

$$X(t) = x + pt - S(t), \quad (1)$$

where  $x$  denotes the initial reserve,  $p$  is the premium rate per unit of time and  $S(t)$  is a stochastic process modeling the amount of cumulative claims up to time  $t$ . Taking  $S$  to be a compound Poisson or compound renewal process yields the Cramér-Lundberg model and the Sparre-Andersen model, respectively. we consider an extensions of classical risk theory towards a multidimensional reserves model (1) where  $X(t)$ ,  $x$ ,  $p$  and  $S(t)$  are vectors, with possible dependence between the components of  $S(t)$ . Indeed, the assumption of independence of risks may easily fail, for example in the case of reinsurance, when incoming claims have an impact on both insuring companies at the same time.

## A general model

Let  $(U_1, U_2)$  be a general Lévy process starting at  $(u_1, u_2)$ .

### Questions:

- The exit time from the positive quadrant:

$$\tau_{\text{or}}(u_1, u_2) = \inf\{t \geq 0 : U_1(t) < 0 \text{ or } U_2(t) < 0\},$$

- The entrance time into the negative quadrant:

$$\tau_{\text{sim}}(u_1, u_2) = \inf\{t \geq 0 : U_1(t) < 0 \ \& \ U_2(t) < 0\}.$$

The associated ultimate/perpetual ruin probabilities will be respectively denoted

$$\psi_{\text{or}}(u_1, u_2) = P(\tau_{\text{or}}(u_1, u_2) < \infty),$$

$$\psi_{\text{sim}}(u_1, u_2) = P(\tau_{\text{sim}}(u_1, u_2) < \infty).$$

Letting  $\tau_i(u_i) = \inf\{t \geq 0 : U_i(t) < 0\}$ ,  $i = 1, 2$ , we also will consider

$$\psi_{\text{and}}(u_1, u_2) = P(\tau_1(u_1) < \infty \text{ and } \tau_2(u_2) < \infty).$$

We study asymptotics and exact expressions.

## **Key-references**

### **Dependent risks/queues**

Ambagaspitiya (1999), Müller (1997) Cai and Li (2005) Chan, Yang and Zhang (2003) Denuit, Genest and Marceau (1999) Dhaene and Goovaerts (1996,1997)  
Lieshout and Mandjes(2006)

### **Local limit theorems and large deviations**

Arfwedson (1955) Höglund (1990)  
Borovkov and Mogulskii (2001) Ignatyuk, Malyshchev and Scherbakov (1994)  
Collamore (1996)

### **One-dimensional ruin**

Asmussen (2000), Rolski et al. (2003), Knessl and Peters (1994) Pervozvansky (1998)

## One-dimensional theory revisited.

Let  $\tau(x)$  denote the first passage time of the level  $-x$  by  $Z$

$$\tau(x) = \inf\{t \geq 0 : x + Z_t < 0\}$$

where and set

$$\psi(x) = P(\tau(x) < \infty)$$

Let  $\kappa$  be the cumulant:  $\kappa(\theta) = \log E[e^{\theta Z_1}]$ .

Let us assume the *Cramér assumption* that there exists a  $\gamma > 0$ , such that  $\kappa(-\gamma) = 0$ .

Under this assumption, it holds that

$$\lim_{x \rightarrow \infty} e^{\gamma x} \psi(x) = C, \quad (2)$$

When  $Z$  is a Cramér-Lundberg process  $C$  is given explicitly by:

$$C = -\kappa'(0)/\kappa'_+(-\gamma). \quad (3)$$

## A general model

Let  $U = (U_1, U_2)$  be a general Lévy process.

Let the cumulant exponent of  $U$  be denoted by  $\kappa$ :

$$e^{\kappa(\theta_1, \theta_2)t} = E[e^{\theta_1(U_1(t) - u_1) + \theta_2(U_2(t) - u_2)}]$$

Assume that  $U$  has light tails, i.e. there exist  $\gamma_i > 0$  s.t.

$$\kappa(-\gamma_1, 0) = \kappa(0, -\gamma_2) = 0.$$

We shall also assume that

$$(-\gamma_1, 0), (0, -\gamma_2) \in \text{int}\{\theta : \kappa(\theta) < \infty\}.$$

Let

$$\psi_i(u_i) = P(\tau_i(u_i) < \infty) = P(\exists t > 0 : U_i(t) < 0)$$

## A general model

If  $K \rightarrow \infty$ , it holds that

$$\psi_{\text{or}}(aK, K) \sim \begin{cases} \psi_1(aK) \sim C_1 e^{-\gamma_1 a K}, & \text{if } \gamma_1 a < \gamma_2 \\ \psi_1(aK) + \psi_2(K) & \text{if } \gamma_1 a = \gamma_2 \\ \psi_2(K) \sim C_2 e^{-\gamma_2 K}, & \text{if } \gamma_1 a > \gamma_2 \end{cases}$$

and

$$\psi_{\text{and}}(aK, K) = o(C_2 e^{-\gamma_2 K} + C_1 e^{-\gamma_1 a K})$$

## Model with co-monotonic dependence

Denote by  $U_i$  the risk process of the  $i$ 'th company

$$U_i(t) := -\delta_i S(t) + c_i t + u_i, \quad i = 1, 2,$$

where  $u_i$  denotes the initial reserve and where

$$S(t) = \sum_{i=1}^{N(t)} \sigma_i, \quad (4)$$

$N(t)$  is a Poisson process with i.i.d. inter-arrival times  $\zeta_i$  and claims  $\sigma_i \geq 0$ .

Assume that the second company, receives less premium per amount paid out, i.e.:

$$p_1 = \frac{c_1}{\delta_1} > \frac{c_2}{\delta_2} = p_2.$$

and that

$$p_i > \rho := E\sigma / E\zeta.$$

## **Geometric considerations**

## Model with co-monotonic dependence

Write

$$X_i(t) = x_i + p_i t - S(t), \quad i = 1, 2,$$

where

$$S(t) = \sum_{i=1}^{N(t)} \sigma_i$$

with  $N$  a Poisson process.

Let  $\kappa_i$  be the cumulant exponent of  $X_i$ :

$$E[e^{\theta X_i(t)}] = e^{\kappa_i(\theta)t}, \quad i = 1, 2. \quad (5)$$

### Questions:

- Explicit solutions of  $\psi_{\text{and}}$ ,  $\psi_{\text{or}}$ ,  $\psi_{\text{sim}}$ .
- What are exact asymptotics of  $\psi_{\text{and}}$ ,  $\psi_{\text{or}}$ ,  $\psi_{\text{sim}}$  if  $x_1, x_2$  go to infinity according to a ray in the positive orthant?

## Boundary crossing problems

Write  $b_1(t) = x_1 + p_1t$ ,  $b_2(t) = x_2 + p_2t$ .

Then the problem can also be reformulated as a one-dimensional first-passage problem over a piecewise linear barrier:

$$\begin{aligned}\psi_{or}(x_1, x_2) &= P(\exists t > 0 : S(t) > \min\{b_1(t), b_2(t)\}) \\ \psi_{sim}(x_1, x_2) &= P(\exists t > 0 : S(t) > \max\{b_1(t), b_2(t)\}) \\ \psi_{and}(x_1, x_2) &= P(\exists t, s > 0 : S(t) > b_1(t), \\ &\quad S(s) > b_2(s))\end{aligned}$$

## Boundary crossing problems

Let

$$T(x_1, x_2) = \frac{x_2 - x_1}{p_1 - p_2}.$$

be the crossing time of  $b_1$  and  $b_2$  and write  $\tau_i(x) = \inf\{t \geq 0 : X_i(t) < 0\}$ .

**Proposition 1** *Suppose that  $x_2 > x_1, p_2 < p_1$ . The ruin probabilities  $\psi_{\text{and}}, \psi_{\text{sim}}$  and  $\psi_{\text{or}}$  are given by:*

$$\begin{aligned} \psi_{\text{sim}}(x_1, x_2) &= P(\tau_2(x_2) \leq T) \\ &+ P(\tau_2(x_2) > T, \inf_{s>T} X_1(s) < 0), \end{aligned}$$

$$\begin{aligned} \psi_{\text{or}}(x_1, x_2) &= P(\tau_1(x_1) \leq T) \\ &+ P(\tau_1(x_1) > T, \inf_{s>T} X_2(s) < 0), \end{aligned}$$

$$\begin{aligned} \psi_{\text{and}}(x_1, x_2) &= P(T < \tau_1(x_1) < \infty) \\ &+ P(\tau_1(x_1) \leq T, \tau_2(x_2) < \infty). \end{aligned}$$

## Exact solution

Let  $q^+(\alpha)$  the largest root of

$$\kappa_1(q) = \alpha$$

and write

$$\bar{\psi}_{or}(x_1, x_2) = 1 - \psi_{or}(x_1, x_2)$$

**Corollary 2** *The double Laplace transform*

$$\tilde{\psi}_{or}(p, q) = \int_0^\infty \int_0^\infty e^{-px_1} e^{-qx_2} \bar{\psi}_{or}(x_1, x_2) dx_1 dx_2$$

is given by

$$\tilde{\psi}_{or}(p, q) = \frac{\kappa_1'(0+) + (p_2 - p_1)}{p(\kappa_1(p + q) - q(p_1 - p_2))} \times \left[ 1 + \frac{p}{q - q^+(q(p_1 - p_2))} \right]$$

**Example.** If  $S$  is compound Poisson with  $\text{Exp}(\mu)$  jumps then  $\tilde{\psi}$  is given by

$$\tilde{\psi}(p, q) = \frac{(\mu + p + q)(p_2 - \rho)}{pp_1(z_1(q) - p)z_2(q)}$$

where  $\rho = \lambda/\mu$  and  $z_1(q) < z_2(q)$  are the roots of  $\kappa_1(p + q) = (p_1 - p_2)q$ .

**Example. Small claims model** If

$$X_i(t) = x_i + p_i t - B(t)$$

where  $B$  is a Brownian motion and  $x_2 > x_1$ , then:

$$\begin{aligned} \psi_{\text{or}}(x_1, x_2) &= 1 - \Phi(a(x_1, p_1)) + e^{-2p_1 x_1} \Phi(a(-x_1, p_1)) \\ &+ e^{-2p_2 x_2} \times [\Phi(a(x_1, p_1 - 2p_2)) \\ &- e^{-2x_1(p_1 - 2p_2)} \Phi(a(-x_1, p_1 - 2p_2))]. \end{aligned}$$

$$\begin{aligned} \psi_{\text{sim}}(x_1, x_2) &= 1 - \Phi(a(x_2, p_2)) + e^{-2p_2 x_2} \Phi(a(-x_2, p_2)) \\ &+ e^{-2p_1 x_1} \times [\Phi(a(x_2, p_2 - 2p_1)) \\ &- e^{-2x_2(p_2 - 2p_1)} \Phi(a(-x_2, p_2 - 2p_1))] \end{aligned}$$

where

$$a(x, p) = [x + pT] / \sqrt{T}$$

and  $\Phi$  denotes the cumulative standard normal distribution function.

## Asymptotic cones

$$\begin{aligned}\mathcal{D}_1 &= \{(x_1, x_2) \in \mathbb{R}_+^2 : x_1 > x_2 s_1\}, \\ \mathcal{D}_2 &= \{(x_1, x_2) \in \mathbb{R}_+^2 : x_1 < x_2 s_2\}\end{aligned}$$

with

$$s_1 = \frac{\kappa'_1(-\gamma_1)}{\kappa'_2(-\gamma_1)}, \quad s_2 = \left( \frac{\kappa'_1(-\gamma_2)}{\kappa'_2(-\gamma_2)} \right)_+$$

and

$$\mathcal{D}_0 = \mathbb{R}_+^2 \setminus [\overline{\mathcal{D}}_1 \cup \overline{\mathcal{D}}_2]$$

where  $\overline{\mathcal{D}}_i$  is the closure of  $\mathcal{D}_i$ , the open cone lying between  $\mathcal{D}_1$  and  $\mathcal{D}_2$ . Write

$$\widehat{\mathcal{D}}_2 = \{(x_1, x_2) : x_1 < x_2 s_3\}$$

where  $s_3 = \kappa'_1(-\widehat{\gamma}_2)/\kappa'_2(-\widehat{\gamma}_2)$  with  $\widehat{\gamma}_2$  is the largest root of

$$\kappa_1(-s) = \kappa_1(-\gamma_2).$$

**Lemma 3** 1.  $\mathcal{D}_i, i = 0, 1, 2$  are disjoint and  $\mathcal{D}_0, \mathcal{D}_1 \neq \emptyset$ .

2.  $\mathcal{D}_1 \subset \mathcal{U} := \{(x_1, x_2) \in \mathbb{R}_+^2 : x_2\gamma_2 < x_1\gamma_1\}$

3.  $\mathcal{D}_2 \neq \emptyset$  iff  $\kappa'_1(-\gamma_2) < 0$ .

**Lemma 4** Writing  $i = 1, 2$

$$T_i = x_i / [-\kappa'_i(-\gamma_i)]$$

and

$$\tilde{T}_i = x_i / [-\kappa'_i(-\gamma_{3-i})]$$

and  $a(x) = \frac{x_1}{x_2}$  (where  $x = (x_1, x_2) \in \mathbb{R}_+^2$ ), the following hold true:

$$\begin{aligned} \mathcal{D}_2 &= \{x \in \mathbb{R}_+^2 : a(x) < s_2\} = \{x \in \mathbb{R}_+^2 : \tilde{T}_1 < T\} \\ &= \{x \in \mathbb{R}_+^2 : T_2 < T\} \end{aligned}$$

$$\mathcal{D}_0 = \{x \in \mathbb{R}_+^2 : T_1 < T < T_2\}$$

$$\begin{aligned} \mathcal{D}_1 &= \{x \in \mathbb{R}_+^2 : s_1 < a(x)\} = \{x \in \mathbb{R}_+^2 : T < \tilde{T}_2\} \\ &= \{x \in \mathbb{R}_+^2 : T < T_1\}. \end{aligned}$$

## Asymptotics

**Notation.** •  $\kappa_2$  denotes the characteristic exponent of  $X_2$  and

$$\kappa^*(v) = \kappa_2^*(v) = \sup_{\beta \in \mathbb{R}} [v\beta - \kappa_2(\beta)]. \quad (6)$$

the convex conjugate of  $\kappa_2$

- Write

$$\underline{\theta} = \inf\{\theta : \kappa_2(\theta) < \infty\} \quad \underline{v} = \lim_{\theta \downarrow \underline{\theta}} \kappa_2'(\theta)$$

- **The adjustment coefficient.**

Assume there exist  $\gamma_i > 0$ , such that

$$\kappa_i(-\gamma_i) = 0$$

and let

$$\gamma(a) = \kappa_2^*(-v_a)/v_a,$$

where

$$v_a := (p_1 - p_2)/(1 - a).$$

## Asymptotics

Set

$$\bar{a} = 1 + \frac{(p_1 - p_2)}{\underline{v}}$$

**Theorem.** Assume that

$$\underline{\theta} < -\gamma_1.$$

For  $a < \bar{a}$ , it holds as  $K \rightarrow \infty$ ,

$$\psi_{\text{sim}}(aK, K) \sim \begin{cases} C_1 e^{-\gamma_1 a K} & \text{if } (aK, K) \in \mathcal{D}_1 \\ (D^\#(v_a) + D'(v_a)) K^{-1/2} e^{-\gamma(a)K} & \text{if } (aK, K) \in \mathcal{D}_0 \\ C_2 e^{-\gamma_2 K} & \text{if } (aK, K) \in \mathcal{D}_2, \end{cases}$$

for some  $D^\#, D' > 0$  and  $C_i = -\kappa'_i(0)/\kappa'_i(-\gamma_i)$ .

## Constants

Denote by  $\theta_v < \theta'_v$  the roots

$$\kappa'_2(\theta_v) = -v$$

and

$$\kappa_2(\theta_v) = \kappa_2(\theta'_v)$$

then

$$D'(w) = \frac{\theta'_w - \theta_w}{|\theta_w \theta'_w|} \frac{\sqrt{w}}{\sqrt{2\pi\kappa''_1(\theta_w)}},$$

$$D^\#(w) = \left[ \frac{1}{\theta_w} - \frac{1}{\theta'_w} + \frac{\kappa'_1(0)}{\kappa_1(\theta'_w)} - \frac{\kappa'_1(0)}{\kappa_1(\theta_w)} \right] \frac{\sqrt{w}}{\sqrt{2\pi\kappa''_1(\theta_w)}}$$

## Asymptotics

Write

$$\widehat{D}_2 = \{(x_1, x_2) : x_1 < x_2 s_3\}$$

where  $s_3 = \kappa'_1(-\widehat{\gamma}_2)/\kappa'_2(-\widehat{\gamma}_2)$  with  $\widehat{\gamma}_2$  is the largest root of

$$\kappa_1(-s) = \kappa_1(-\gamma_2).$$

**Theorem.** For  $v_a < -\underline{v}$  it holds that, as  $K \rightarrow \infty$ ,

$$\psi_{\text{and}}(aK, K) \sim \begin{cases} C_1 e^{-\gamma_1 a K} & \text{if } (aK, K) \in \mathcal{D}_1 \\ (D^*(v_a) - D^\dagger(v_a)) K^{-1/2} e^{-\gamma(a)K} & \text{if } (aK, K) \in \widehat{\mathcal{D}}_0 \\ \widehat{C}_2 e^{-(a\widehat{\gamma}_2 + (1-a)\gamma_2)K} & \text{if } (aK, K) \in \widehat{\mathcal{D}}_2, \end{cases}$$

for some  $D^*, D^\dagger > 0$  where

$$\widehat{C}_2 = C_2 \kappa'_1(-\gamma_2)/\kappa'_1(-\widehat{\gamma}_2).$$

## Constants.

Denote by  $\theta_v^* > \theta_v$  the root of

$$\kappa_1(\theta_v) = \kappa_1(\theta_v^*)$$

then  $D^*(w)$  and  $D^\dagger(w)$  are respectively given by with  $\theta'_w$  replaced by  $\theta_w^*$  and  $\kappa_1$  replaced by  $\kappa_2$ .

## One-dimensional theory revisited.

Write  $f(x) \sim h(x)$  ( $x \rightarrow \infty$ ) if  $\lim f/h(x) = 1$  and

$$w(x, t) = \psi(x) - \psi(x, t) = P(t < \tau(x) < \infty).$$

### Höglund-Arfwedson Theorem

Assume that either  $\kappa'_+(0) < 0$  or that the Cramér assumption holds and write  $\zeta = \min\{\theta : \kappa(\theta) = 0\}$ . If  $0 < v < -\underline{v}$  and  $x, t \rightarrow \infty$  such that  $x/t = v$ , it holds that

$$\psi(x, t) \sim \begin{cases} Ce^{-x\zeta} & \text{if } x/t < -\kappa'_+(-\zeta) \\ |D(v)|t^{-1/2}e^{-t\kappa^*(-v)} & \text{if } x/t > -\kappa'_+(-\zeta), \end{cases}$$
$$w(x, t) \sim \begin{cases} |D(v)|t^{-1/2}e^{-t\kappa^*(-v)} & \text{if } x/t < -\kappa'_+(-\zeta) \\ Ce^{-x\zeta} & \text{if } x/t > -\kappa'_+(-\zeta), \end{cases}$$

with

$$D(v) = c(v) \cdot \frac{1}{\sqrt{2\pi\kappa''(\theta_v)}}, \quad c(v) = \frac{\theta'_v - \theta_v}{\theta_v\theta'_v},$$

## Asymptotic Limit laws

The following result shows existence of and identifies the limit laws of the Cramér-Lundberg process  $X(t)$  conditioned on  $t < \tau$  and on  $t > \tau$ :

**Theorem** Assume that  $\underline{\theta} < 0$  and  $\kappa'(0) < 0$ .

(i) If  $0 < v < -\kappa'(0)$ ,

$$\bar{\Psi}_v = \lim_{x,t \rightarrow \infty, x=tv} P(X(t) \in \cdot \mid \tau(x) > t),,$$

in the sense of weak convergence, where

$$\bar{\Psi}_v(dy) = c(v)^{-1} [e^{-\theta'_v y} - e^{-\theta_v y}] \mathbf{1}_{(0,\infty)}(y) dy.,$$

(ii) If  $-\kappa'(0) < v < -\underline{v}$ ,

$$\Psi_v = \lim_{x,t \rightarrow \infty, x=tv} P(X(t) \in \cdot \mid \tau(x) < t)$$

in the sense of weak convergence, where

$$\Psi_v(dy) = |c(v)|^{-1} [e^{-\theta'_v y} \mathbf{1}_{(0,\infty)}(y) + e^{-\theta_v y} \mathbf{1}_{(-\infty,0)}(y)] d$$

## Two term expansions

We write  $f \approx g + h$  if  $\lim_{x \rightarrow \infty} (f - g)/h(x) = \lim_{x \rightarrow \infty} (f - h)/g(x) = 1$ .

**Proposition** If  $x_1, x_2 \rightarrow \infty$  such that  $x_2/T(x_1, x_2) = v$ ,  $v \neq -\kappa'_2(-\gamma_i)$   $i = 1, 2$ , it holds that

$$\begin{aligned} \psi_{\text{sim}}(x_1, x_2) &\approx \psi_2(x_2, T) \\ &+ \tilde{C}_1(v) e^{-\gamma_1 x_1} \bar{\psi}_2^{(-\gamma_1)}(x_2, T), \end{aligned}$$

for some  $C_1(v)$ . If  $S$  is a compound Poisson process, it holds for  $i = 1, 2$ ,

$$\tilde{C}_i(v) = \begin{cases} C_i & \text{if } -\kappa'_2(-\gamma_i) < v \\ c_{3-i}(v, \gamma_i)^{-1} [\psi_i^*(\theta_v) - \psi_i^*(\eta_v^{(i)})] & \text{if } 0 < v < -\kappa'_2(-\gamma_i), \end{cases}$$

and  $\eta_v^{(1)} = \theta'_v$ ,  $\eta_v^{(2)} = \theta_v^*$ .

## Renewal model – Asymptotics

$$S(t) = \sum_{i=1}^{N(t)} \sigma_i, \quad (7)$$

where  $N(t)$  is a renewal process with i.i.d. inter-arrival times  $\zeta_i$ , and the claims  $\sigma_i$  are i.i.d. positive random variables independent of  $N(t)$ .

### Cramer/light tail assumptions

Assume that there exist constants  $\gamma_i > 0$  ( $i = 1, 2$ ) such that

$$E[e^{\gamma_i(\sigma_1 - p_i \zeta_1)}] = 1. \quad (8)$$

Moreover, assume that, for some  $\epsilon > 0$ ,

$$E[e^{(\gamma_1 + \epsilon)\sigma_1}] < \infty. \quad (9)$$

**Theorem 5** *Let  $a < 1$ . Under the Cramer conditions it holds that, as  $K \rightarrow \infty$ ,*

$$\begin{aligned} \psi_{\text{or}}(aK, K) &\sim C_2 e^{-\gamma_2 K} + C_1 e^{-\gamma_1 aK}, \\ \psi_{\text{and}}(aK, K) &= o(C_2 e^{-\gamma_2 K} + C_1 e^{-\gamma_1 aK}). \end{aligned}$$